

# Design and Analysis of Occupancy Studies

## Part 4

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## Multiple-States Single-Season Occupancy Studies

# Multiple States; Single-Season

## Objectives:

- Estimate “gross” occupancy rates.
- Partition occupancy into a number of hierarchical states (e.g. non-breeders vs. breeders)

# Multiple States; Single-Season - Sampling Protocol

## Sampling Protocol:

- Landscape divided (artificially or naturally) into  $S$  patches or cells or SITES.
- Select  $s \ll S$  sites at random (all sites have equal probability of selection).
- Visit each site  $K$  times .
- Record detection or not detection of species in site  $i$  in survey  $k$ .
- If occupancy detected, classify (with detection error) state of occupancy (e.g. adults seen, vs. adults seen with young.)  
[Most software can only deal with 2 states.]

## Note:

- Two levels of false negatives, i.e. not detecting a species does not imply that site was unoccupied, and given that see an animal, no detection of breeding does not imply that breeding did not take place.
- States can be classified in a hierarchy.

# Multiple States; Single-Season - Dynamics

Data: Detection/Encounter histories with ., 0, 1, and 2.

- 1021 Confirmed breeding status in survey 3.
- 0101 Occupied, but breeding status is unknown.
- 0000 Unsure if unoccupied or occupied.

Parameters:

- $\psi_i^1$  Prob( site  $i$  is occupied regardless of state).
- $\psi_i^2$  Prob( site  $i$  in state 2 (typically breeders) given that site is occupied.
- $p_{ij}^1$  Prob( detection in site  $i$  in survey  $j$  given that in state 1, non-breeder).
- $p_{ij}^2$  Prob( detection in site  $i$  in survey  $j$  given that in state 2, breeder).
- $\delta_{ij}$  Prob( identify as breeder (state 2) in site  $i$  in survey  $j$  given that site is occupied, in state 2, and species detected).

Then  $\text{Prob}(\text{occupied in state 2}) = \psi^1 \psi^2$ .

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Prob(1021) =

$$\psi_i^1 \psi_i^2 p_{i1}^2 (1 - \delta_{i1}) (1 - p_{i2}^2) p_{i3}^2 \delta_{i3} p_{i4}^2 (1 - \delta_{i4})$$

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Prob(0101) =

$$\psi_i^1(1 - \psi_i^2) [(1 - p_{i1}^1)p_{i2}^1(1 - p_{i3}^1)p_{i4}^1] + \\ \psi_i^1\psi_i^2 [(1 - p_{i1}^2)p_{i2}^2(1 - \delta_{i2})(1 - p_{i3}^2)p_{i4}^2(1 - \delta_{i4})]$$

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Prob(0000) = Long and messy expression!



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- ① Occupancy STATE of sites is constant during season (closure).
- ② Probability of occupancy in each state ( $\psi^1, \psi^2$ ) is equal across all sites (homogeneity) unless modified by covariates.
- ③ Probability of detection ( $p^1, p^2$ ) given occupancy is equal across all sites (homogeneity) unless modified by covariates.
- ④ Detection in each survey of a site is independent of those on other surveys
- ⑤ Detections at each site are independent
- ⑥ No false positives.
- ⑦ **NEW** Reproduction has occurred at start of “season” and evidence of reproduction continues for entire season (i.e. reproduction doesn’t start in survey 5 or 10, and you must know that breeding took place even after chicks fledged).
- ⑧ **NEW** Evidence of states independent across surveys. Not valid to return to a new nest once chick found in subsequent surveys as no longer independent. Truncate record at confirmed reproduction and fill remainder with missing values.

# Multiple States; Single-Season - Biological Hypotheses

- ① Do different habitats have different occupancy in state 1 and state 2?
- ② Is detection in state 1 and state 2 the same?
- ③ Sink and Source habitat identification.

# Multiple States; Single-Season - Example

Nichols et al (2007) Ecology 88:1395-1400 looked at breeding and non-breeding California Spotted Owls in  $s = 54$  sites over  $k = 5$  surveys.

Hand-fed mice confirmed occupancy; breeding status only confirmed when owl took mouse to nest.

- . = site not visited.
- 0 = owl not detected.
- 1 = owl detected, but no detection of young.
- 2 = owl detected, along with evidence of young.

# Multiple States; Single-Season - Example

## Using MARK

Locate the data file, and start MARK in the usual fashion:

Enter Specifications for MARK Analysis

Select Data Type

- ☐ Live Recaptures (CJS)
- ☐ Dead Recoveries (Seber)
- ☐ Both (Burnham)
- ☐ Known Fates
- ☐ Closed Captures
- ☐ BTQ Ring Recoveries
- ☐ Robust Design
- ☐ Both (Barker)
- ☐ Multi-state Recaptures only
- ☐ Dead Recoveries (Brownie et al.)
- ☐ Jolly-Seber
- ☐ Pradel Models Including Robust D
- ☐ Barker Robust Design
- ☐ POPAN
- ☐ VPA -- Virtual Population Analysis
- ☐ Multi-state -- Live and Dead Enc.
- ☐ Multi-state -- Live and Dead Enc.
- ☒ Occupancy Estimation
- ☐ 2-Species Occupancy Estimation
- ☐ Robust Design Occupancy

Title for this set of data:

**Occupancy Data Types -- Pick One**

- Occupancy Estimation
- Occupancy Heterogeneity Estimation
- Occupancy Estimation Royle/Nichols Poisson
- Occupancy Estimation Royle/Nichols Negative Binomial
- Occupancy Estimation Royle Poisson Counts
- Occupancy Estimation Royle Negative Binomial Counts
- Multiple-state Occupancy Estimation**
- Multi-scale Occupancy Estimation

Help Cancel OK

States: 2 Enter State Names

Mixtures: 2

Default Time Intervals Used

Default Group Labels Used

Default Ind. Cov. Names Used

Default State Names Used

# Multiple States; Single-Season - Example

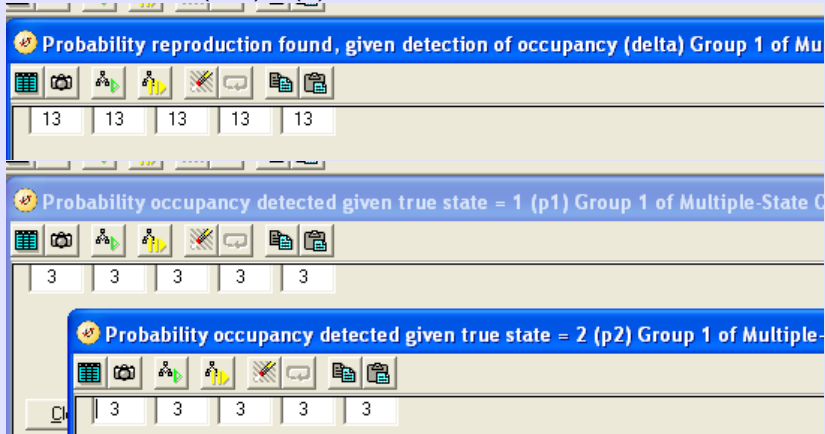
Biological models:

- $p_{it}^s = p$ , or  $p_{it}^s = p^s$  or  $p_{it}^s = p_t$  or  $p_{it}^s = p_{it}^s$
- $\delta_{it} = \delta$  or  $\delta_{it} = \delta_{it}$  or  $\delta_{it} = \delta^{12}$  or  $\delta^{345}$  based on it being “hard” to detect reproduction early and “easier” to detect reproduction later in season.

Start by fitting the the model  $\psi^1, \psi^2, p(*, *), \delta(*)$

# Multiple States; Single-Season - Example

Model  $\psi^1, \psi^2, p(*, *), \delta(*)$



# Multiple States; Single-Season - Example

Model  $\psi^1, \psi^2, p(*, *), \delta(*)$  results. Interpret these.

Model	AICc	Delta AICc	AICc Weight	Model Likelihood	No. Par.	Deviance	-2Log(L)
{psi1 psi2 p(*,*) delta(*)}	336.1948	0.0000	1.00000	1.0000	4	0.0000	327.3785

California Spotted Owl - Multistate Model				
Real Function Parameters of {psi1 psi2 p(*,*) delta(*)}				
Parameter	Estimate	Standard Error	95% Confidence Interval	
			Lower	Upper
1:Psi1	0.9752276	0.0377781	0.6474845	0.9988162
2:Psi2	0.5542889	0.1124421	0.3375910	0.7521436
3:p1	0.7448009	0.0363100	0.6674445	0.8093044
4:Delta	0.4229782	0.0843411	0.2713339	0.5906726

California Spotted Owl - Multistate Model				
Estimates of Derived Parameters				
Estimates of Psi1*Psi2 {psi1 psi2 p(*,*) delta(*)}				
Group	Psi1*Psi2-hat	Standard Error	95% Confidence Interval	
			Lower	Upper
1	0.5405578	0.1116381	0.3277341	0.7395495

# Multiple States; Single-Season - Example

Other models to fit:

- $\psi^1, \psi^2, p(*, t), \delta(*)$
- $\psi^1, \psi^2, p(s, *), \delta(*)$
- $\psi^1, \psi^2, p(s, t), \delta(*)$
- $\psi^1, \psi^2, p(*, *), \delta(12, 345)$
- $\psi^1, \psi^2, p(*, t), \delta(12, 345)$

Results:

Model	AICc	Delta AICc	AICc Weight	Model Likelihood	No. Par.	Deviance	-2Log(L)
{psi1 psi2 p(*,t) delta(12,345)}	283.4459	0.0000	0.95457	1.0000	9	0.0000	261.3550
{psi1 psi2 p(*,s) delta(12,345)}	289.5992	6.1533	0.04402	0.0461	5	0.0000	278.3492
{psi1 psi2 p(s,t) delta(12,345)}	296.4865	13.0406	0.00141	0.0015	14	0.0000	257.7172
{psi1 psi2 p(*,t) delta(*)}	329.5843	46.1384	0.00000	0.0000	8	0.0000	310.3843
{psi1 psi2 p(*,s) delta(*)}	336.1948	52.7489	0.00000	0.0000	4	0.0000	327.3785
{psi1 psi2 p(s,*) delta(*)}	338.6127	55.1668	0.00000	0.0000	5	0.0000	327.3627
{psi1 psi2 p(*,t) delta(*)}	341.3668	57.9209	0.00000	0.0000	13	0.0000	306.2668



# Multiple States; Single-Season - Example

## Results:

California spotted owl - Multistate Model				
Real Function Parameters of {psi1 psi2 p(*,t) delta(12,345)}				
Parameter	Estimate	Standard Error	95% Confidence Interval	
			Lower	Upper
1:Psi1	0.9789489	0.0323620	0.6817037	0.9990106
2:Psi2	0.4495324	0.0801697	0.3020490	0.6064560
3:p1	0.5790117	0.0821281	0.4154070	0.7269281
4:p1	0.8237920	0.0677819	0.6518159	0.9211063
5:p1	0.9223752	0.0429091	0.7859100	0.9746595
6:p1	0.6970106	0.0771335	0.5292754	0.8247649
7:p1	0.6196330	0.1122511	0.3904316	0.8055704
8:Delta	0.2935477E-016	0.1555513E-008	-0.3048805E-008	0.3048805E-008
9:Delta	0.8683310	0.0729096	0.6539372	0.9583604

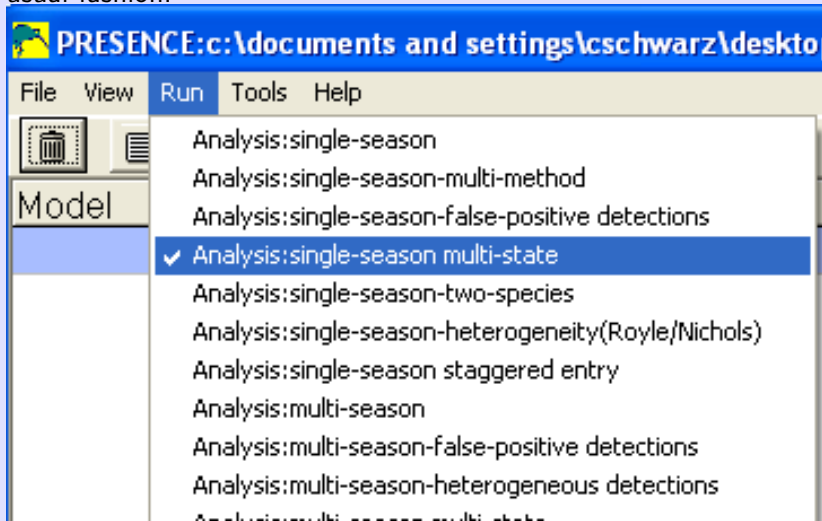
  

California spotted owl - Multistate Model				
Estimates of Derived Parameters				
Estimates of Psi1*Psi2 {psi1 psi2 p(*,t) delta(12,345)}				
Group	Psi1*Psi2-hat	Standard Error	95% Confidence Interval	
			Lower	Upper
1	0.4400692	0.0798189	0.2940524	0.5972503

Unfortunately, GOF and  $\hat{c}$  not yet implemented in MARK.

## Using PRESENCE

Locate the data file, start PRESENCE, and enter the data in the usual fashion:



# Multiple States; Single-Season - Example

Biological models:

- $p_{it}^s = p$ , or  $p_{it}^s = p^s$  or  $p_{it}^s = p_t$  or  $p_{it}^s = p_{it}^s$
- $\delta_{it} = \delta$  or  $\delta_{it} = \delta_{it}$  or  $\delta_{it} = \delta^{12}$  or  $\delta^{345}$  based on it being “hard” to detect reproduction early and “easier” to detect reproduction later in season.

Start by fitting the the model  $\psi^1, \psi^2, p(*, *), \delta(*)$ .

Note: PRESENCE uses  $R$  for  $\psi^2$ .

# Multiple States; Single-Season - Example

Model  $\psi^1, \psi^2, p(*, *), \delta(*)$

Design Matrix - Single			Design Matrix - Single		Design Matrix - Single	
File Init Retrieve model			File Init Retrieve model		File Init Retrieve model	
Occ/Brd	Detection		Occ/Brd	Detection	Occ/Brd	Detection
-	a1	a2	-	b1	-	c1
psi	1	0	p1(1)	1	dlta(1)	1
R	0	1	p1(2)	1	dlta(2)	1
			p1(3)	1	dlta(3)	1
			p1(4)	1	dlta(4)	1
			p1(5)	1	dlta(5)	1
			p2(1)	1		
			p2(2)	1		
			p2(3)	1		
			p2(4)	1		
			p2(5)	1		

# Multiple States; Single-Season - Example

Model  $\psi^1, \psi^2, p(*, *), \delta(*)$  results. Interpret these.

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Individual site estimates of <psi>				
	Site	estimate	Std.err	95% conf. interval
psi	1 site 1	: 0.9752	0.0378	0.6475 - 0.9988

Individual site estimates of <R>				
	Site	estimate	Std.err	95% conf. interval
R	1 site 1	: 0.5543	0.1124	0.3376 - 0.7521

Individual site estimates of <p1(1)>				
	Site	estimate	Std.err	95% conf. interval
p1(1)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093
p1(2)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093
p1(3)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093
p1(4)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093
p1(5)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093

Individual site estimates of <p2(1)>				
	Site	estimate	Std.err	95% conf. interval
p2(1)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093
p2(2)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093
p2(3)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093
p2(4)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093
p2(5)	1 site 1	: 0.7448	0.0363	0.6674 - 0.8093

Individual site estimates of <d1ta(1)>				
	Site	estimate	Std.err	95% conf. interval
d1ta(1)	1 site 1	: 0.4230	0.0843	0.2713 - 0.5907
d1ta(2)	1 site 1	: 0.4230	0.0843	0.2713 - 0.5907
d1ta(3)	1 site 1	: 0.4230	0.0843	0.2713 - 0.5907
d1ta(4)	1 site 1	: 0.4230	0.0843	0.2713 - 0.5907
d1ta(5)	1 site 1	: 0.4230	0.0843	0.2713 - 0.5907

# Multiple States; Single-Season - Example

Other models to fit:

- $\psi^1, \psi^2, p(*, t), \delta(*)$
- $\psi^1, \psi^2, p(s, *), \delta(*)$
- $\psi^1, \psi^2, p(s, t), \delta(*)$
- $\psi^1, \psi^2, p(*, *), \delta(12, 345)$
- $\psi^1, \psi^2, p(*, t), \delta(12, 345)$

Results:

Model	AIC	deltaAIC	AIC wgt	Model Likel	no.Par.	-2*LogLike
psi,R. p(*,t) delta(12,345)	279.36	0.00	0.9890	1.0000	9	261.36
psi,R. p(*,*) delta(12,345)	288.35	8.99	0.0110	0.0112	5	278.35
psi,R. p(*,t) delta(*)	326.38	47.02	0.0000	0.0000	8	310.38
psi,R. p(s,t) delta(*)	332.27	52.91	0.0000	0.0000	13	306.27
psi,R. p(*,*) delta(*)	335.38	56.02	0.0000	0.0000	4	327.38
psi,R. p(s,*) delta(*)	337.36	58.00	0.0000	0.0000	5	327.36

# Multiple States; Single-Season - Example

## Results:

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=====
```

Individual site estimates of <psi>				
	site	estimate	std.err	95% conf. interval
psi	1 site 1	: 0.9789	0.0324	0.6817 - 0.9990

Individual site estimates of <R>				
	site	estimate	std.err	95% conf. interval
R	1 site 1	: 0.4495	0.0802	0.3020 - 0.6065

Individual site estimates of <p1(1)>				
	site	estimate	std.err	95% conf. interval
p1(1)	1 site 1	: 0.5790	0.0821	0.4154 - 0.7269
p1(2)	1 site 1	: 0.8238	0.0678	0.6518 - 0.9211
p1(3)	1 site 1	: 0.9224	0.0429	0.7859 - 0.9747
p1(4)	1 site 1	: 0.6970	0.0771	0.5293 - 0.8248
p1(5)	1 site 1	: 0.6196	0.1123	0.3904 - 0.8056

Individual site estimates of <p2(1)>				
	site	estimate	std.err	95% conf. interval
p2(1)	1 site 1	: 0.5790	0.0821	0.4154 - 0.7269
p2(2)	1 site 1	: 0.8238	0.0678	0.6518 - 0.9211
p2(3)	1 site 1	: 0.9224	0.0429	0.7859 - 0.9747
p2(4)	1 site 1	: 0.6970	0.0771	0.5293 - 0.8248
p2(5)	1 site 1	: 0.6196	0.1123	0.3904 - 0.8056

Individual site estimates of <d1ta(1)>				
	site	estimate	std.err	95% conf. interval
d1ta(1)	1 site 1	: 0.0000	0.0000	0.0000 - 1.0000
d1ta(2)	1 site 1	: 0.0000	0.0000	0.0000 - 1.0000
d1ta(3)	1 site 1	: 0.8683	0.0729	0.6539 - 0.9584
d1ta(4)	1 site 1	: 0.8683	0.0729	0.6539 - 0.9584
d1ta(5)	1 site 1	: 0.8683	0.0729	0.6539 - 0.9584

## Multiple-States Single-Season Occupancy Studies

Using *RPresence* software.



# Multiple States; Single-Season

## Objectives:

- Estimate “gross” occupancy rates.
- Partition occupancy into a number of hierarchical states (e.g. non-breeders vs. breeders)

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[Most software can only deal with 2 states.]

## Note:

- Two levels of false negatives, i.e. not detecting a species does not imply that site was unoccupied, and given that see an animal, no detection of breeding does not imply that breeding did not take place.
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Then  $\text{Prob}(\text{occupied in state 2}) = \psi r$ .

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Prob(1021) =

$$\psi_i r_i p_{i1}^2 (1 - \delta_{i1}) (1 - p_{i2}^2) p_{i3}^2 \delta_{i3} p_{i4}^2 (1 - \delta_{i4})$$

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- $\delta_{ij}$  Prob( identify as breeder (state 2) in site  $i$  in survey  $j$  given that site is occupied, in state 2, and species detected).

Prob(0101) =

$$\psi_i(1 - r_i) [(1 - p_{i1}^1)p_{i2}^1(1 - p_{i3}^1)p_{i4}^1] + \\ \psi_i r_i [(1 - p_{i1}^2)p_{i2}^2(1 - \delta_{i2})(1 - p_{i3}^2)p_{i4}^2(1 - \delta_{i4})]$$

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- ④ Detection in each survey of a site is independent of those on other surveys
- ⑤ Detections at each site are independent
- ⑥ No false positives.
- ⑦ **NEW** Reproduction has occurred at start of “season” and evidence of reproduction continues for entire season (i.e. reproduction doesn’t start in survey 5 or 10, and you must know that breeding took place even after chicks fledged).
- ⑧ **NEW** Evidence of states independent across surveys. Not valid to return to a new nest once chick found in subsequent surveys as no longer independent. Truncate record at confirmed reproduction and fill remainder with missing values.

# Multiple States; Single-Season - Biological Hypotheses

- ① Do different habitats have different occupancy in state 1 and state 2?
- ② Is detection in state 1 and state 2 the same?
- ③ Sink and Source habitat identification.



## Multiple States; Single-Season - Example - *RPresence*

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- 0 = owl not detected.
- 1 = owl detected, but no detection of young.
- 2 = owl detected, along with evidence of young.

Sites selected based on assumed occupancy so  $\psi$  is not relevant. Visits divided in two periods (early (visits 1,2), and late (visits 3,4,5)).

Using *RPresence*

# Multiple States; Single-Season - Example - *RPresence*

Read in the history data

```
1 occ.data <- readxl::read_excel("CalSpottedOwl.xls",  
2                               sheet="RawData",  
3                               skip=10, na='.')  
4 head(occ.data)  
5  
6 input.history <- occ.data[,-1] #drop first column  
7  
8 Nsites <- nrow(input.history)  
9 Nvisits<- ncol(input.history)
```

Define the site-time covariates. These are in visit major order.

```
1 period <- data.frame(  
2     Site=rep(1:Nsites, Nvisits),  
3     Visit=rep(1:Nvisits, each=Nsites),  
4     Per  =rep(c("E","E","L","L","L"), each=Nsites))
```

# Multiple States; Single-Season - Example - *RPresence*

Create the \*.pao object

```
1 owl.pao <- createPao(input.history,  
2                       survcov=period,  
3                       title="owl multi state single season"  
4 summary(owl.pao)
```

Naive occ=0.8703704

naiveR =0.3518519

nunits	nsurveys	nseasons
"54"	"5"	"1"
nsurveyseason	nmethods	nunitcov
"5"	"1"	"1"
nsurvcov		
"4"		

unit covariates : TEMP

survey covariates: SURVEY Site Visit Per

Notice naive estimates of occupancy.

Notice strange number of unit covariates?

Biological models:

- $p_{it}^s = p$ , or  $p_{it}^s = p^s$  or  $p_{it}^s = p_t$  or  $p_{it}^s = p_{it}^s$
- $\delta_{it} = \delta$  or  $\delta_{it} = \delta_{it}$  or  $\delta_{it} = \delta^{12}$  or  $\delta^{345}$  based on it being “hard” to detect reproduction early and “easier” to detect reproduction later in season.

Start by fitting the the model  $\psi, r, p1(*), p2(*), \delta(*)$

# Multiple States; Single-Season - Example - *RPresence*

Model  $\psi, r, p1(*) = p2(*), \delta(*)$

```
1 mod1 <- occMod(model=list(  
2     psi~1, # occupancy regardless of state  
3     r~1,   # occupancy in state 2 / occupied  
4     p~1,   # detection in state i  
5     delta~1), # identified as state 2 if detected (and in s  
6     data=owl.pao,  
7     type="do.ms.2")  
8 summary(mod1)
```

Model name=psi()r()p()delta()

AIC=335.3784

-2\*log-likelihood=327.3784

num. par=4

What are the parameters?

## Multiple States; Single-Season - Example - *RPresence*

Model  $\psi, r, p1(*) = p2(*), \delta(*)$

```
> mod1$real$psi[1,] # prob site is occupied
```

	est	se	lower	upper
--	-----	----	-------	-------

unit1	0.9752	0.0378	0.6475	0.9988
-------	--------	--------	--------	--------

```
> mod1$real$r[1,] # prob in state 2 | occupied
```

	est	se	lower	upper
--	-----	----	-------	-------

unit1	0.5543	0.1124	0.3376	0.7521
-------	--------	--------	--------	--------

$\psi$  is not of interest (see previous) but compare to naive estimate.

$r$  estimates proportion of breeders - compared to naive estimate.



## Multiple States; Single-Season - Example - *RPresence*

Model  $\psi, r, p1(*) = p2(*), \delta(*)$

	est	se	lower	upper
unit1_1-1	0.7448	0.0363	0.6674	0.8093
unit1_1-2	0.7448	0.0363	0.6674	0.8093
...				

```
> mod1$real$p2[seq(1,by=Nsites, length.out=Nvisits),] # p2
```

	est	se	lower	upper
unit1_1-1	0.7448	0.0363	0.6674	0.8093
unit1_1-2	0.7448	0.0363	0.6674	0.8093
...				

```
> mod1$real$delta[seq(1,by=Nsites, length.out=Nvisits),] # delta
```

	est	se	lower	upper
unit1_1-1	0.423	0.0843	0.2713	0.5907
unit1_1-2	0.423	0.0843	0.2713	0.5907
...				

Less than 50% chance of identifying a breeding pair if actually a breeding pair!

Other models to fit:

- $\psi, r, p1(t) = p2(t), \delta(*)$
- $\psi, r, p1(*), p2(*), \delta(*)$
- $\psi, r, p1(*) = p2(*), \delta(Per)$
- $\psi, r, p1(t) = p2(t), \delta(Per)$
- $\psi, r, p1(*), p2(*), \delta(Per)$

What do these models mean. Check the output to confirm.

Caution with model that involve  $\delta(Per)$  as the detection probability in period 1 is 0 which causes numerical problems.

# Multiple States; Single-Season - Example - *RPresence*

AICc table:

	Model	DAIC	wgt	npar	neg2ll	warr
1	psi()r()p(SURVEY)delta(Per)	0.00	0.9848	9	261.36	
2	psi()r()p()delta(Per)	8.99	0.0110	5	278.35	
3	psi()r()p(STATE)delta(Per)	10.88	0.0043	6	278.24	
4	psi()r()p(SURVEY)delta()	47.03	0.0000	8	310.38	
5	psi()r()p()delta()	56.02	0.0000	4	327.38	
6	psi()r()p(STATE)delta()	58.01	0.0000	5	327.36	

# Multiple States; Single-Season - Example - *RPresence*

Model averaged estimates

```
> RPresence::modAvg(results, param="psi")[1,]  
      est      se lower_0.95 upper_0.95  
unit1 0.9788445 0.03249207   0.681173   0.999003
```

```
> RPresence::modAvg(results, param="r") [1,]  
      est      se lower_0.95 upper_0.95
```

$\psi$  not useful (why?).

Interpret estimate of  $r$ .

# Multiple States; Single-Season - Example - *RPresence*

Model averaged estimates

```
> RPresence::modAvg(results, param="p1")[seq(1,by=Nsites, 1
```

	est	se	lower_0.95	upper_0.95
--	-----	----	------------	------------

unit1_1-1	0.5814721	0.08402199	0.4138997	0.7321398
-----------	-----------	------------	-----------	-----------

...

```
> RPresence::modAvg(results, param="p2")[seq(1,by=Nsites, 1
```

	est	se	lower_0.95	upper_0.95
--	-----	----	------------	------------

unit1_1-1	0.5815911	0.08426813	0.4135243	0.7326347
-----------	-----------	------------	-----------	-----------

...

```
> RPresence::modAvg(results, param="delta")[seq(1,by=Nsites
```

	est	se	lower_0.95	upper_0.95
--	-----	----	------------	------------

unit1_1-1	2.593953e-11	NaN	NaN	NaN
-----------	--------------	-----	-----	-----

unit1_1-2	2.593953e-11	NaN	NaN	NaN
-----------	--------------	-----	-----	-----

unit1_1-3	8.683107e-01	0.07269637	0.6547659	0.9581997
-----------	--------------	------------	-----------	-----------

unit1_1-4	8.683107e-01	0.07269637	0.6547659	0.9581997
-----------	--------------	------------	-----------	-----------

unit1_1-5	8.683107e-01	0.07269637	0.6547659	0.9581997
-----------	--------------	------------	-----------	-----------

Interpret estimate of  $\delta$ .

Unfortunately, GOF and  $\hat{c}$  not yet implemented in *RPresence* but you could do a bootstrap after defining your appropriate statistic.

## Multiple States; Single-Season - Exercise

Coosa bass collected via electrofishing from four 50-80 m sections in streams at 54 sites in the Upper Coosa River basin. The objective was to evaluate the effect of streamflow variability on Coosa bass reproduction. States are:

- species not detected (state = 0),
- adult detected (state = 1),
- YOY present (state = 2).

There are 2 covariates stream link magnitude (standardized), and CV of stream flow during summer.

# Multiple States; Single-Season - Exercise

Models to fit:

- $\psi^1, \psi^2, p(*, *), \delta(*)$
- $\psi^1(L), \psi^2(L), p(*, *)(L), \delta(*) (L)$ , i.e. as a function of length.
- $\psi^1(CV), \psi^2(CV), p(*, *)(CV), \delta(*) (CV)$ , i.e. as a function of variation in flow

Hints:

- Does “time” have a meaning here?
- Don’t forget that a logistic regressions has an intercept and a slope.



# Multiple States; Single-Season - Exercise

## PRESENCE results:

Model	AIC	deltaAIC	AIC wgt	Model Likel	no.Par.	-2*LogLike
psi,R, p1(*)=p2(*) (L) ,deta(*) (L)	360.74	0.00	0.9999	1.0000	8	344.74
psi,R, p1(*)=p2(*) (CV) ,deta(*) (CV)	380.02	19.28	0.0001	0.0001	8	364.02
psi,R, p1(*)=,p2(*) , delta(*)	389.20	28.46	0.0000	0.0000	4	381.20

# Multiple States; Single-Season - Exercise

MARK results:

Model	AICc	Delta AICc	AICc Weight	Model Likelihood	No. Par.	Deviance	-2Log(L)
{psi1(L), psi2(L) p("")(L) delta(")(L)}	363.9405	0.0000	0.99993	1.0000	8	344.7405	344.7405
{psi1(CV), psi2(CV) p("")(CV) delta(")(CV)}	383.2208	19.2803	0.00007	0.0001	8	364.0208	364.0208
{psi1, psi2 p("")( ) delta(")( )}	390.0119	26.0714	0.00000	0.0000	4	381.1956	381.1956

# Multiple States; Single-Season - Exercise

## MARK results:

Coose Bass				
LOGIT Link Function Parameters of {psi1(L), psi2(L)}			p(*,*)(L) 95% Confidence Interval	delta(*)(L)} 95% Confidence Interval
Parameter	Beta	Standard Error	Lower	Upper
1:	3.7267939	1.4143066	0.9547530	6.4988349
2:	3.5339919	1.2414517	1.1007466	5.9672372
3:	1.4197249	0.9709173	-0.4832731	3.3227228
4:	-0.5373002	0.7437811	-1.9951112	0.9205108
5:	1.5582874	0.2434488	1.0811277	2.0354470
6:	-0.2513000	0.2129819	-0.6687446	0.1661446
7:	-0.5722787	0.3185632	-1.1966626	0.0521052
8:	0.1457854	0.2874052	-0.4175289	0.7090997

Coose Bass				
Real Function Parameters of {psi1(L), psi2(L)}			p(*,*)(L)	delta(*)(L)}
Following estimates based on unstandardized individual covariate values:				
Variable	Value			
STDLINK	0.1146296			
CVFLOW	1.4740741			

Parameter	Estimate	Standard Error	95% Confidence Interval	
			Lower	Upper
1:Psi1	0.9842012	0.0240583	0.7501328	0.9992270
2:Psi2	0.7954563	0.1476977	0.3962592	0.9584072
3:p1	0.8219304	0.0336287	0.7463313	0.8786621
4:Delta	0.3645737	0.0711546	0.2391065	0.5116103

## Multiple-State Single-season Summary

# Multiple States; Single-Season - Summary

Similar to previous methods + :

- Key parameters are
  - $\psi^1$  = prob of occupancy regardless of state 1 (non-breeder) or state 2 (breeder)
  - $\psi^2$  = prob of state 2 (breeder) given that species present on site.
  - Must adjust for two types of false negative – non-detect of occupancy; non-detect of breeding if occupancy confirmed.
  - Planning studies will require much thought. Use GENPRES in a similar fashion as in previous examples.

## EXTENSIONS:

- Multiple seasons – colonization and extinction could depend on occupancy and breeding success.